

# Econometrics II TA Session \*

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## 1 Review of Truncated Regression Model

### 1.1 Truncated normal variable

Let's start from a random variable  $Z \sim N(0, 1)$ , and the distribution function is

$$\phi(z) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{z^2}{2}\right),$$

$$\Phi(z) = \int_{-\infty}^z \phi(z) dz.$$

Accordingly, the distribution function of  $X \sim N(\mu, \sigma^2)$  can be derived as

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) = \frac{1}{\sigma} \phi\left(\frac{x-\mu}{\sigma}\right).$$

Pay attention to the cumulative distribution function of  $x$ , it should be

$$F_x(a) = \int_{-\infty}^a f(x) dx = \Phi\left(\frac{a-\mu}{\sigma}\right).$$

Now we assume an interval  $(K_1, K_2)$ , where  $K_1$  and  $K_2$  are known. The conditional density  $X|K_1 < X < K_2$  (using the definition of conditional density) is given by

$$f(x|K_1 < x < K_2) = \frac{f(x)}{\int_{K_1}^{K_2} f(x) dx} = \frac{\frac{1}{\sigma} \phi\left(\frac{x-\mu}{\sigma}\right)}{\Phi\left(\frac{K_2-\mu}{\sigma}\right) - \Phi\left(\frac{K_1-\mu}{\sigma}\right)}$$

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\*The codes are cited from documents by Hiroki Kato.

Then

$$\begin{aligned} E[x|K_1 < x < K_2] &= \int_{K_1}^{K_2} xf(x|K_1 < x < K_2)dx \\ &= \sigma \frac{\phi\left(\frac{K_1 - \mu}{\sigma}\right) - \phi\left(\frac{K_2 - \mu}{\sigma}\right)}{\Phi\left(\frac{K_2 - \mu}{\sigma}\right) - \Phi\left(\frac{K_1 - \mu}{\sigma}\right)} + \mu \end{aligned}$$

## 1.2 Truncated regression model

We assume that  $y_i, \mathbf{x}_i$  are continuous random variables and the selection rule is

$$y_i = \mathbf{x}_i\boldsymbol{\beta} + u_i \quad \text{if } a_1 < y < a_2$$

where  $u_i \sim N(0, \sigma^2)$ . with  $a_1, a_2$  known constant values. If  $y_i$  falls into  $(a_1, a_2)$ , then we observe both  $y_i, \mathbf{x}_i$ . If not, we do not observe  $y_i$  or  $\mathbf{x}_i$ . Therefore, groups of data, which exists in reality but does not fit the selection rule, is not included or observed.

The probability density function of  $y_i$  conditional on  $\mathbf{x}_i$  and  $a_1 < y < a_2$  should be

$$p_\theta(y_i|\mathbf{x}_i, a_1 < y < a_2) = \frac{f(y_i|\mathbf{x}_i)}{P(a_1 < y < a_2|\mathbf{x}_i)}$$

where  $\theta = (\boldsymbol{\beta}, \sigma^2)'$ . Considering the distributional assumption and taking variable transformation, the conditional distribution of  $y_i$  is given by

$$p(y_i|\mathbf{x}_i) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2} \left(\frac{y_i - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}\right)^2\right) = \frac{1}{\sigma} \phi\left(\frac{y_i - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}\right),$$

where  $\phi(\cdot)$  is the standard normal density function. Moreover, the probability of observation  $y_i > 0$  is given by

$$\begin{aligned} P(a_1 < y < a_2|\mathbf{x}_i) &= P(a_1 < \mathbf{x}_i\boldsymbol{\beta} + u_i < a_2|\mathbf{x}_i) \\ &= P\left(\frac{a_1 - \mathbf{x}_i\boldsymbol{\beta}}{\sigma} < \frac{u_i}{\sigma} < \frac{a_2 - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}|\mathbf{x}_i\right) \\ &= \Phi\left(\frac{a_2 - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}\right) - \Phi\left(\frac{a_1 - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}\right) \end{aligned}$$

where  $\Phi(\cdot)$  is the standard normal cumulative density function.

Thus, the log-likelihood function is supposed to be

$$M_n(\theta) = \sum_{i=1}^n \log \left( \frac{1}{\sigma} \frac{\phi\left(\frac{y_i - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}\right)}{\Phi\left(\frac{a_2 - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}\right) - \Phi\left(\frac{a_1 - \mathbf{x}_i\boldsymbol{\beta}}{\sigma}\right)} \right)$$